

# Estimating the small-scale galaxy correlation function

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## 1 The projected correlation function (Davis & Peebles, 1983)

### 1.1 Notation

One can measure the two-point correlation function in redshift space  $\xi^{\text{obs}}(r_p, \pi)$  as a function of two variables,  $\pi$ , the radial separation, and  $r_p$ , the transverse separation. The observables are redshifts (due to the Hubble flow and intrinsic velocities) and angular positions on the sky. For  $z \ll 1$ , we have (with  $c = 1$ ) :

$$\pi = (z_1 - z_2)/H_0 \quad , \quad r_p \equiv (z_1 + z_2)/H_0 \times \tan(\theta_{12}/2) \quad (1)$$

The redshift separation between two galaxies is defined as

$$s \equiv (z_1^2 + z_2^2 - 2z_1z_2 \cos(\theta_{12}))^{1/2} / H_0 \quad (2)$$

It approximates the true separation  $r$  when peculiar velocities are negligible.

On small scales, where  $\Delta z \ll 1$ ,  $\theta_{12} \ll 1$ , we have  $s \approx (\pi^2 + r_p^2)^{1/2}$ .

By small scales, we refer to  $\sim 1 - 10$  Mpc.

### 1.2 Getting rid of the peculiar velocities

For a given true radial separation  $y$ , and total separation  $r = (r_p^2 + y^2)^{1/2}$ , the measured radial separation  $\pi$  can be decomposed as follows :

$$\pi = y - y h(r) + v_{\text{pec}}/H_0 \quad (3)$$

The first term is due to the Hubble expansion. The second term comes from bulk motions ( $h(r) = 0$  if the cluster expands with the general expansion on the scale  $r$ ,  $h(r) = 1$  if the cluster is Virialized,  $h(r) > 1$  if the cluster is collapsing). The third term comes from the random peculiar velocities in the cluster along the line of sight.

Given a probability distribution for the peculiar velocities  $f(v_{\text{pec}})$ , the observed correlation function is then obtained by the following convolution of the real space correlation function  $\xi(r)$  with  $f$  :

$$1 + \xi^{\text{obs}}(r_p, \pi) = \int_{-\infty}^{+\infty} [1 + \xi(r)] H_0 f(v_{\text{pec}}) dy \quad , \quad r = (r_p^2 + y^2)^{1/2} \quad , \quad v_{\text{pec}} = H_0(\pi - y + y h(r)) \quad (4)$$

Note that typically, people use  $f(v) \propto e^{-\sqrt{2}|v|/\sigma}$ .

From this we obtain the projected correlation function :

$$w_p(r_p) \equiv \int_{-\infty}^{+\infty} d\pi \xi^{\text{obs}}(r_p, \pi) = \int_{-\infty}^{+\infty} dy \xi \left( (r_p^2 + y^2)^{1/2} \right) \quad (5)$$

With this procedure, one therefore gets rid of the effect of the peculiar velocities.

This equation can be inverted to recover  $\xi(r)$  from the measured  $w_p(r_p)$  :

$$\xi(r) = -\frac{1}{\pi} \int_r^{+\infty} dr_p w'(r_p) (r_p^2 - r^2)^{-1/2} \quad (6)$$

## 2 A more robust estimator (Padmanabhan et al., 2007)

### 2.1 Theory

In practise the integral in  $w_p$  cannot be extended to infinity and needs to be cut at a cutoff scale  $\pi_{\max} \gg r_p$ . This renders the estimator sensitive to poorly measured long wavelength modes. Padmanabhan et al. (2007) suggest a new estimator  $\omega$ , using high-pass filtering to remove the problematic long wavelength modes. They define

$$\omega(r_s) \equiv 2\pi \int_0^{r_s} dr_p r_p G(r_p, r_s) w_p(r_p) \quad , \quad w_p(r_p) \equiv \int_{-\pi_{\max}}^{\pi_{\max}} d\pi \xi^{\text{obs}}(r_p, \pi) \quad , \quad (7)$$

where the filter  $G$  has a characteristic scale  $r_s$  :  $G(r_p, r_s) = r_s^{-3} g(r_p/r_s)$ , and is ‘‘compensated’’ , i.e. s.t.

$$\int_0^{r_s} dr_p r_p G(r_p, r_s) = 0 \quad . \quad (8)$$

This ensures that the slowly varying modes are canceled out.

$\omega(r_s)$  can also be related to the real space correlation function  $\xi(r)$ . Note that for a power-law correlation function, one obtains  $\omega(r_s) \propto \xi(r_s)$ .

### 2.2 Filter $G(r_p, r_s)$

The family of filters  $G(r_p, r_s) = r_s^{-3} g(r_p/r_s)$  is considered, with

$$g(x) = x^{2\alpha} (1 - x^2)^\beta (c - x^2) \quad , \quad (9)$$

where the constant  $c$  is chosen such that the integral of  $r_p G(r_p, r_s)$  vanishes.

In practise, they adopt  $(\alpha, \beta) = (2, 2)$ .

### 2.3 Practical calculation

$\xi^{\text{obs}}$  can be estimated via

$$\xi^{\text{obs}}(r_p, \pi) = \frac{DD(r_p, \pi)}{RR(r_p, \pi)} - 1 \quad (10)$$

Therefore, the estimator  $\omega$  is given by

$$\omega(r_s) = 2\pi \int_0^{r_s} dr_p r_p G(r_p, r_s) \int_{-\pi_{\max}}^{\pi_{\max}} d\pi \frac{DD(r_p, \pi)}{RR(r_p, \pi)} \quad (11)$$

The constant disappears because of the insensitivity of  $G$  to constant changes.

For small bins  $\Delta\pi$  ,  $\Delta r_p$ , one can write

$$RR(r_p, \pi) = 2\pi r_p^2 \bar{n} N \Phi(r_p, \pi) \Delta \ln r_p \Delta \pi \quad , \quad (12)$$

where

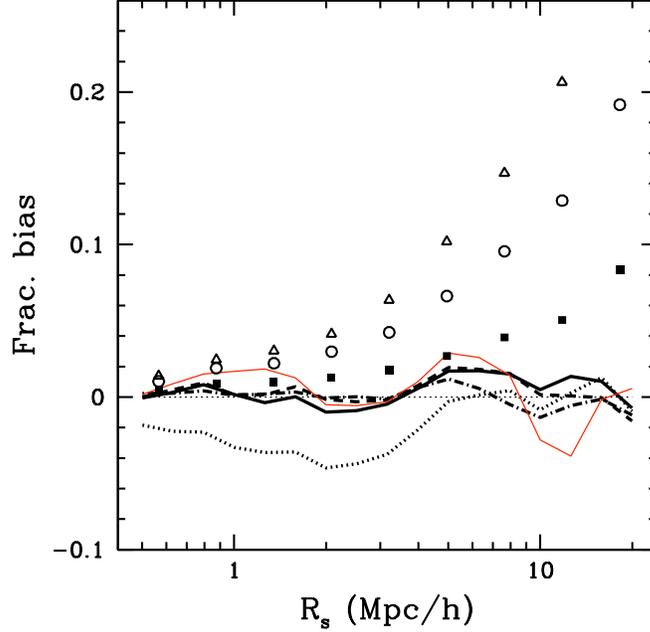
$$\Phi(r_p, \pi) \equiv \frac{\int S(\mathbf{x}_1) S(\mathbf{x}_1 + \mathbf{s}) d^3 \mathbf{x}_1}{\int S(\mathbf{x}_1) d^3 \mathbf{x}_1} \quad , \quad \mathbf{s} \equiv \mathbf{r}_p + \pi \mathbf{e}_\pi \quad (13)$$

If one takes a binning small enough that  $DD$  is either 0 or 1, then one has

$$\omega(r_s) = \sum_{i \in DD, r_p, i < r_s, |\pi_i| < \pi_{\max}} \frac{G(r_p, i, r_s)}{\bar{n} N \Phi(r_p, i, \pi_i)} \quad , \quad (14)$$

i.e. the integral is transformed into a simple Riemann sum.

*Note* : the authors claim that this estimator is insensitive to the integral constraint. My understanding is that the integral constraint is an uncertainty in the multiplicative factor in  $\xi$ . The high-pass filtering removes additive constants in  $\xi$ . I am not quite sure how it works then.



**Figure 4.** The fractional bias in  $\omega$  ( $\text{bias} = (\omega - \omega_{\text{ref}})/\omega_{\text{ref}}$ ) obtained from the redshift space distribution of galaxies in the Millennium simulation for different  $Z_{\text{max}}$  (in  $\text{Mpc}/h$ ) – 20 (dotted), 40 (thick solid), 60 (dashed), and 80 (dot-dashed). The true value ( $\omega_{\text{ref}}$ ) is assumed to be obtained by integrating to a  $Z_{\text{max}} = 100\text{Mpc}/h$ ; the figure shows that  $\omega$  has converged to better than 2% by  $Z_{\text{max}} = 40\text{Mpc}/h$ . The thin solid (red) line plots  $\omega$  obtained by integrating over the real space correlation function. The fluctuations in  $\Delta\omega$  are consistent with measurement noise. For comparison, the points show the analogous bias terms for  $w_p$ , with triangles, circles, and squares corresponding to  $Z_{\text{max}} = 40, 60,$  and  $80\text{Mpc}/h$  respectively. For a fair comparison with  $\omega$ , we plot the bias in  $w_p$  at  $R_s/2$ , which roughly corresponds to the central scale probed by  $\omega$ . Note that the biases in  $w_p$  are significantly larger than those for  $\omega$ .

## 2.4 Testing the filter with the Millenium simulation

One of the main advantages of  $\omega$  versus  $w_p$  is that it converges faster when one increases  $\pi_{\text{max}}$ . It therefore completes the partial removal of the peculiar velocity distortions originally intended by  $w_p$ .  $\omega$  is converged to better than 2% by  $\pi_{\text{max}} = 40h^{-1} \text{Mpc}$ . This can be seen their figure 4 reproduced above (where  $Z_{\text{max}} = \pi_{\text{max}}$ ).

## References

- Davis & Peebles, 1983, ApJ, 267, 465  
 Padmanabhan, White & Eisenstein, 2007, MNRAS, 376, 1702